

7/1/22

Agenda: Exam answer, and intro to Prob 2.

Answers [Brief]

Part (A)

- (1) If X, Y are the results upon rolling a dice twice, then X, Y obviously have the same distribution i.e. $U[1, 6]$, whose expectations are $\frac{7}{2}$. So $E[X] = E[Y] = \frac{7}{2}$.
- (2) The only conditions that are required to preserve pmfs are non-negativity and $\sum_{x \in \Omega} f(x) = 1$. So, squaring is unlikely to preserve pmfs. On the other hand, $\frac{p+a}{2}$ is a pmf, and $\sum_{x \in \Omega} p(x/2) = 1/2$ so it can never be a pmf.
- (3) If $A \subseteq B$, then $A \cap B = A, A \cup B = B$. Therefore, $A \cap B^c = \emptyset$. We easily see that $P(B|A) = P(A^c|B^c) = 1$ and $P(A|B^c) = 0 = P(B^c|A)$. That $P(A|B) = 1 - P(A^c|B)$ is a swift consequence of the finite additivity of $P(\cdot|B)$.
- (4) It's about the quantities $\lim_{x \rightarrow -\infty} F(x), \lim_{x \rightarrow \infty} F(x)$ because the candidates are all continuous. But: $\lim_{x \rightarrow -\infty} e^{-x} = \infty, \lim_{x \rightarrow -\infty} 1 - \frac{e^{-x}}{2} = -\infty, \lim_{x \rightarrow \infty} 1 - 2e^{-x} = 1$. and $\lim_{x \rightarrow \infty} 1/x^2 = 0$: So NONE of them are DFs.
- (5) According to Defn 25.7 and Thm 25.10, two PDFs represent the same random variable even if they differ at one point. Now, the operations $x \mapsto x+1$ and $x \mapsto x^2$ are continuous: so $Y = X+1, Z = X^2$ are also continuous r.v.s. Having said that, it is not necessary that Z needs to have a PDF with $f_Z(z) = 0 \nmid z \leq 0$, even if $P(Z \leq 0) = 0$. For example, take the two PDFs:
 $f_X = \mathbb{1}_{\{0 < x < 1\}}$ $f'_X = \mathbb{1}_{\{0 < x < 1\}} + \mathbb{1}_{\{x=2\}}$
 They're both pdfs: look at 25.7. But both are a pdf for the uniform $[0, 1]$ random variable! So, $f_Z(z) = 0 \nmid z \leq 0$ is false $f_Y(y) = 0 \nmid y \leq 1$ is false anyway, let $X \sim U[-1, 0]$ e.g.

Part B

- (1) The key point: $A_{k_1} \cap \dots \cap A_{k_m} = A_{\text{lcm}(k_1, \dots, k_m)}$. Indeed, j is a multiple of $k_1, \dots, k_m$ if and only if it is a multiple of $\text{lcm}(k_1, \dots, k_m)$. Thus, combining with $P(A_i) = \frac{1}{i}$
 $P(A_{k_1} \cap \dots \cap A_{k_m}) = P(A_{\text{lcm}(k_1, \dots, k_m)}) = \frac{1}{\text{lcm}(k_1, \dots, k_m)}$
 We now see that if $\{k_1, k_2, k_3\}$ are pairwise coprime, then $\forall S \subseteq \{1, 2, 3\}$ we have
 $\text{lcm}(\{k_s : s \in S\}) = \prod_{s \in S} k_s$
 and therefore, as a consequence, $P(\cap_{s \in S} A_{k_s}) = \prod_{s \in S} P(A_{k_s})$.
 So $\{A_{k_i} : i=1, 2, 3\}$ is an independent colln. However, if k_1, k_2 are not coprime, then $A_{k_1} \cap A_{k_2} \neq A_{k_1 k_2}$ so clearly A_{k_1} is not independent of A_{k_2} .

- (2) An exercise in cond'l probability. Define the events:
 $E_i = \{i^{\text{th}} \text{ ball is red}\}$
 $E_{ij} = \{i^{\text{th}} \& j^{\text{th}} \text{ balls are red}\}, i < j$
 $E_{ijk} = \{i^{\text{th}}, j^{\text{th}} \& k^{\text{th}} \text{ balls are red}\}, i < j < k$

Then $X_{ijk} = \mathbb{1}_{E_{ijk}}$, so $E[X_{ijk}] = P(E_{ijk})$ and $\text{Var}(X_{ijk}) = P(E_{ijk})P(E_{ijk}^c)$.

- (1) $P(E_i) = r/btr$. Claim: $P(E_i) = r/btr \forall i$.
 pf: $P(E_i) = P(E_i | E_1)P(E_1) + P(E_i | E_1^c)P(E_1^c)$
 $\downarrow \quad \quad \quad \downarrow \quad \quad \quad \downarrow$
 $\frac{r+c}{btr+c} \quad \frac{r}{btr} \quad \frac{r}{btr+c} \quad \frac{b}{btr}$
 $= r/btr$

- (2) $\forall i > 1, P(E_i) = P(E_i | E_1)P(E_1) = \frac{r(r+c)}{btr(btr+c)}$
 Claim: $P(E_{ij}) = \frac{r(r+c)}{btr(btr+c)} \forall i < j$.
 pf: $P(E_{ij}) = P(E_j | E_i)P(E_i)$.

$P(E_j | E_i) = P(E_j \cap E_i | E_i) + P(E_j \cap E_i^c | E_i)$
 $= P(E_j | E_i)P(E_i | E_i) + P(E_j | E_i^c)P(E_i^c | E_i)$

Use induction hypothesis along with reasoning of (1) to conclude.

- (3) For this, see that

$$P(E_{ijk}) = \frac{r(r+c)(r+c+c)}{btr(btr+c)(btr+c)}$$

Then use:

$$P(E_{ijk} | E_i) = P(E_{jk} | E_i)P(E_i | E_i) + P(E_{jk} | E_i^c)P(E_i^c | E_i)$$

along with ind. hyp. & reasoning of parts (1), (2) to conclude that

$$P(E_{ijk}) = \frac{r(r+c)(r+c+c)}{btr(btr+c)(btr+c)}$$

- (3) An exercise in binomials.

- (a) $\{\text{System working}\} = \{\exists S \subseteq [5], |S| \geq 3, \text{ exactly the components with labels in } S \text{ are working}\}$
 $\Rightarrow P(\{\text{System working}\}) = \sum_{S \subseteq [5], |S| \geq 3} \prod_{i \in S} p_i \prod_{i \notin S} (1-p_i)$

- (b) Obvious: $P(\text{System working}) = P(\text{Bin}(n, p) \geq k)$

- (c) $P(L, n \text{ not working, System working}) = P(\text{Bin}(n-2, p) \geq k)(1-p)^2$
 Answer = $1 - \frac{P(\text{Bin}(n, p) \geq k)}{P(\text{Bin}(n, p) \geq k)}$

- (4) (a) Obviously, $P(|X| \leq 0) = 0$. Let $x > 0$. Then
 $P(|X| \leq x) = P(x \leq X \leq x) = \int_0^x \frac{1}{2b} e^{-t/b} dt$
 $\stackrel{\text{cov}}{=} 2 \int_0^x \frac{1}{2b} e^{-t/b} dt = [-e^{-t/b}]_0^x = [-e^{-t/b}]_0^x = 1 - e^{-x/b}$.
 So:
 $f_X(x) = \begin{cases} 0 & x \leq 0 \\ 1 - e^{-x/b} & x > 0 \end{cases}$ & $f_X(x) = \begin{cases} 0 & x \leq 0 \\ \frac{e^{-x/b}}{b} & x > 0 \end{cases}$
 Note: $|X| \sim \text{Exp}(1/b)$.

(b) $E[X^k] = \int x^k e^{-x/b} / 2b dx = 0$ if x is odd, as the integrand is an odd fn. $\int_{-\infty}^{\infty} x^k e^{-x/b} dx$.
 If k is even, $E[X^k] = E[|X|^k] = \int_0^{\infty} \frac{x^k}{b} e^{-x/b} dx$.

Let $I_n = \int_0^{\infty} \frac{x^n}{b} e^{-x/b} dx$; so $I_0 = 1$ and

$$\int_0^{\infty} x^n \left[\frac{e^{-x/b}}{b} \right] dx = \int_0^{\infty} x^n e^{-x/b} dx = [-x^n e^{-x/b}]_0^{\infty} - \int_0^{\infty} n x^{n-1} e^{-x/b} dx = nb \int_0^{\infty} x^{n-1} e^{-x/b} dx$$

$$\Rightarrow I_n = (nb) I_{n-1} = n(n-1)b^2 I_{n-2} = \dots = b^n n! I_0 = b^n n!$$

$$\Rightarrow E[X^k] = \begin{cases} 0 & \text{odd} \\ k! b^n & \text{even} \end{cases}$$

[Can use result on Gamma distribution's moments as well]

- (5) $f(x) = \alpha x^{-\alpha-1} \mathbb{1}_{x \geq 1}$.

- (a) Clearly, $P(X^2 \leq 0) = 0$. Let $x > 0$.

$$P(X^2 \leq x) = P(X \leq \sqrt{x}) = \int_1^{\sqrt{x}} \alpha t^{-\alpha-1} dt = [-t^{-\alpha}]_1^{\sqrt{x}} = 1 - x^{-\alpha/2}$$

$$\text{So, } F_{X^2}(x) = \begin{cases} 0, & x \leq 1 \\ 1 - x^{-\alpha/2}, & x > 1 \end{cases} \Rightarrow f_{X^2}(x) = \begin{cases} 0 & x \leq 1 \\ \frac{\alpha}{2} x^{-\alpha/2-1} & x > 1 \end{cases}$$

- (b) Clearly, $P(X^{-1} \notin (0, 1)) = 0$. So let $a \in (0, 1)$.

$$P(X^{-1} \leq a) = P(X \geq a^{-1}) = \int_{a^{-1}}^{\infty} \alpha t^{-\alpha-1} dt = [-t^{-\alpha}]_{a^{-1}}^{\infty} = a^{\alpha}$$

$$\text{So, } F_{X^{-1}}(x) = \begin{cases} 0 & x \leq 0 \\ 1 & x \geq 1 \\ x^{\alpha} & x \in (0, 1) \end{cases} \Rightarrow f_{X^{-1}}(x) = \begin{cases} 0 & x \notin (0, 1) \\ \alpha x^{\alpha-1} & x \in (0, 1) \end{cases}$$

- (c) Note that

$$E[(X^{-1})^k] = \alpha \int_0^1 x^{k+\alpha-1} dx = \alpha \left[\frac{x^{k+\alpha}}{k+\alpha} \right]_0^1 = \frac{\alpha}{k+\alpha} \quad k+\alpha \neq 0$$

$$= \alpha [\ln x]_0^1 \quad k+\alpha = 0$$

Now it is clear that $E[(X^{-1})^k] = \begin{cases} \frac{\alpha}{k+\alpha} & k+\alpha > 0 \\ \pm \infty & k+\alpha \leq 0 \end{cases}$

So, $E[(X^{-1})^k] < \infty \Leftrightarrow k+\alpha > 0$ and $= \frac{\alpha}{k+\alpha}$ in that case.

Probability 2

A list of what you learn in prob 2, with a one-line "why".

- Review of prob 1.
- Multi-variable probability: To extend all prob. 1 notions to multi-variable situation.
- Marginals: Connecting m.v. probability with single variable probability.
- Generating Functions: Linking discrete combinatorial quantities with continuous quantities such as differentiable functions (generally, power series)
- Conditional Prob: Extension to multi-variable situation.
- Notions of Convergence: Understanding various ways in which two random variables can be "close" to each other, or can be compared with each other.
- Central Limit Theorem and laws of large Numbers: Understanding (only proving the weak LLN's finite variance case) these fundamental results in probability: that speak about the accuracy and error of estimation by repeated averaging.

Good references

- Generating Functionology: Herbert Wilf
- Kai Lai Chung's "Probability"
- Klenke's Probability Theory: Read this to understand prob. 3 as well.

Good Luck!